

1 **Sharing work and food within the household:**
2 **Intra-couple time allocation effects on**
3 **nutritional outcomes in rural Telangana, India**

4
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6 *Feminist Economist – In press*

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8
9 **Abstract**

10 *In the context of the ongoing rural transformation in many countries, women's*
11 *opportunities for economic participation are expanding. However, there is a limited*
12 *understanding of how policy interventions can support rural households to adapt to the*
13 *increasing opportunity cost of women's time in household activities. This paper presents*
14 *empirical evidence on the relationship between couple interdependencies in time use and*
15 *nutritional outcomes in rural Telangana, India. We used innovative datasets that*
16 *combined accelerometer-based physical activity data, time use, food intake, and socio-*
17 *demographic data – within the Actor-Partner Interdependence Model (APIM)*
18 *framework. We find that differences in time allocation patterns between spouses in a*
19 *household affect individual nutritional outcomes; when the male spouse allocates more*
20 *time to economic activities, it tends to reduce the adequacy of the female spouse's energy*
21 *intake, and conversely, when the female spouse allocates more time to domestic activities,*
22 *it tends to reduce the male spouse's energy intake adequacy.*

- 1 **Keywords:** Intra-couple time allocation; nutritional outcomes; physical activity;
- 2 actor-partner interdependence model; Telangana, India.
- 3 **JEL codes :** J22, Q12, D13

1 Introduction

2 Malnutrition continues to be a development challenge in many low- and middle-income
3 countries (LMICs), where around 185 million people cannot afford sufficient daily energy
4 intake at an average cost of \$0.79 (FAO et al., 2020). In the past decades, many agricultural
5 and development interventions aimed at enhancing, diversifying, and substituting livelihood
6 means have targeted women based on the central role that they play in ensuring household
7 nutrition (FAO, 2011; Fiorella et al., 2016; Haddad et al., 1997). Even though interventions
8 targeted at women can lead to greater participation in economic activities, increased
9 productivity, and higher household incomes, it is not certain that nutritional outcomes will
10 improve. This disconnect may arise because women in male-headed households still lack the
11 capacity to influence household decision-making regarding the allocation of increased income
12 and the use of their own time. (Kadiyala et al., 2014).

13 Building on the intrahousehold resource allocation literature, this paper presents empirical
14 evidence on spousal interdependencies in time use and nutritional outcomes by investigating
15 own and partner effects¹ of intrahousehold work division on nutritional outcomes among rural
16 households in rural Telangana, India. We hypothesize that the time allocation of male and
17 female spouses and the interdependencies between both individuals are consequential for
18 nutritional outcomes. Inequity in intrahousehold work division has been linked to
19 malnutrition, as women disproportionately bear household domestic work (including child-
20 care) in addition to economic labour (Gillespie et al., 2012). Recent debates have suggested
21 that sharing domestic responsibilities with men will reduce the burden borne by women and
22 potentially improve women's well-being (Asadullah & Kambhampati, 2021; Madzorera &

¹ Own effects capture intra-individual outcomes (e.g., the effect of women's time use in a specific activity on their own nutritional outcome) while partner effects, interpersonal outcomes (e.g., the effect of women's time use in a specific activity on the spouse's nutritional outcome).

1 Fawzi, 2020; Rao & S. Raju, 2020), yet successful policy action to redefine men's household
2 role will require understanding the well-being outcomes for both women and men.

3 Two are the major motivations for this paper. First, there is a research gap in the assessment
4 of an individuals' partner effects, despite the extensive literature on household behaviour².
5 We assess partner effects in this study based on the premise that within households, couples
6 share work and they also share food, especially in rural agricultural contexts where production
7 and consumption decisions are interwoven (C. R. Doss & Quisumbing, 2020; Folbre, 1986;
8 Singh et al., 1986). Even when couples adopt separate economic production spheres, it can be
9 expected that they share some production and consumption between them. Second, there is
10 thin evidence on how spouses in rural agricultural households in LMICs are adapting to rural
11 transformation. Agriculture remains a major contributing sector to the rural economy in terms
12 of employment and income, as increases in agricultural labour productivity and
13 mechanization of farm activities are leading rural households to diversify into the rural non-
14 farm sector (IFAD, 2016). Rural non-farm employment participation, including the time
15 allocated to such activities is dominated by men (Binswanger-Mkhize, 2013; IFAD, 2016),
16 but there are more opportunities for women to participate in economic activities outside of the
17 home as a result of better education, changing socio-cultural norms and improvements in
18 rural-urban transportation linkages (Binswanger-Mkhize, 2013; Ohlan, 2016). This form of
19 rural transformation increases women's opportunity cost of time spent on food preparation

² A large body of literature looked at intrahousehold dynamics to understand the paradigms of household behaviour (see Fafchamps and Quisumbing (2007) for a review). The economic theory of household behaviour proposed by the unitary model aggregates utility of household members (Becker, 1981). Collective models contrarily posit independence in individual preferences and in the process of decision making. They argue that intrahousehold allocation is guided by bargaining even when couples cooperate (Apps & Rees, 1997; Chiappori, 1992; Lundberg & Pollak, 1993; McElroy & Horney, 1981). The drawback to the collective approach is when bargaining for food and other goods like healthcare and leisure is bounded by cultural norms, the approach produces outcomes akin to the unitary model (Agarwal, 1997; Duflo & Udry, 2004).

1 and care activities. The ensuing time reallocation can have implications for household
2 nutritional outcomes (Da Corta & Venkateshwarlu, 1999).

3 Our analysis adopts the Actor Partner Interdependence Model (APIM) framework to capture
4 partner and own effects. APIM is a model of interdependency between individuals (e.g.
5 women and men) in a dyadic relationship (e.g. wife and husband). It postulates that own and
6 partner's characteristics simultaneously influence the outcomes of both individuals (Cook &
7 Kenny, 2005). The analysis allows to capture the effects of one person's characteristics on
8 own outcomes and on the other person's outcome (partner effect). In studying bidirectional
9 effects on two individuals, APIM presents a straightforward transition from economic theories
10 which have either considered households as unified in their interests and preferences or
11 treated individuals as independent decision-making units. The identification of women's and
12 men's time use and their implications on their own and partner's nutritional outcomes is the
13 contribution of this paper to the intrahousehold allocation literature. We examine the
14 following research question: Among couples in farming households in rural Telangana
15 (India), how does the distribution of time spent on economic, domestic, and leisure activities
16 impact not only their own nutritional outcomes but also those of their partners?

17 **2 Literature review**

18 This literature review explores time allocation and nutritional outcomes in rural agricultural
19 settings, first distilling evidence on the effects of women's time allocation on own nutritional
20 outcomes and then focussing on the effects of intrahousehold time allocation on the
21 nutritional outcomes of other family member(s).

1 **2.1 Women’s time allocation and nutritional outcomes**

2 As a result of the ongoing rural transformations, there has been an increase in the number of
3 women active in agriculture, including in the time women allocate to agricultural activities
4 across all regions in LMICs – a trend known as “feminization of agriculture”(Asadullah &
5 Kambhampati, 2021; FAO, 2011). Data collected from the rural areas of Telangana in India
6 shows that women now spend on average, an additional two hours per day in agricultural
7 activities than men and perform male-associated tasks such as land clearing, irrigation and
8 plant protection on the farm (Padmaja et al., 2019). Conversely, male time commitment to
9 farm work is on a downward trend due to the mechanization of male-dominated tasks and the
10 result of male out-migration from rural areas (Padmaja et al., 2019). These changes in male
11 and female time allocation are expected to have consequences for nutritional outcomes, yet
12 empirical evidence is still very limited.

13 Women’s time use tends to have a strong effect on their own well-being outcomes, however,
14 the direction of the effect is not univocal (Ghosh & Bharati, 2005; Johnston et al., 2018; Ruel
15 et al., 2018). Komatsu et al., (2018) found that women’s agricultural time use is associated
16 with a reduction in the consumption of diverse diets among women in Mozambique. They
17 reported better nutritional outcomes among individuals in poor farming households. Other
18 studies reported a negative association between agricultural time use and nutritional outcomes
19 as well. The limited available evidence on this topic suggests that the ability of women to
20 translate agricultural time allocation into desirable nutritional outcomes is mediated by
21 diverse factors. Ghosh & Bharati (2005) found that the effect of agricultural time allocation
22 on body mass index is mediated by socio-demographic factors, although women in paid
23 agricultural work experienced better nutrition than unpaid working women. Also examining
24 the differentiating effects of paid and unpaid work on household nutrition among women in

1 five Indian states, Sangwan & Kumar (2021) found that women in paid farm work have better
2 nutrition compared with peers in non-paid work – as a result of the increase in bargaining
3 power emanating from women’s labour force participation. Further, the effects of time use on
4 the nutritional wellbeing of women and men vary across agricultural seasons as the energy
5 demand of work is highest during land maintenance and harvest seasons (Picchioni et al.,
6 2020; Rao & S. Raju, 2020; Srinivasan et al., 2020). This seasonality effect is intensified
7 among individuals in non-mechanized farming households (Daum et al., 2019; Komatsu et al.,
8 2019) and the landless (Vemireddy & Pingali, 2021). In their review, Johnston et al., (2018)
9 concluded that increased time allocated to agriculture, and the resulting nutritional outcomes
10 will depend on how different individuals in an agricultural household respond to the changes
11 in time use.

12 In addition to women’s time allocation to agricultural activities, women (and girls)
13 disproportionately perform more than three-quarters of household domestic and care work
14 (Jacques Charmes, 2019). However, the evidence linking participation in domestic work and
15 wellbeing in rural areas is very sparse. Often, time use in domestic activity is explained in the
16 context of trade-offs with agricultural and childcare activities, but not how it directly relates
17 to wellbeing. Desai & Jain (1994) argue that domestic work reduces women’s available time
18 to both childcare and economic activities; to the extent that domestic work can be a greater
19 obstacle than childcare to female labour force participation. A multi-country study across
20 Asia and Africa on women’s time use and dietary diversity found that time spent cooking is
21 positively associated with women’s dietary diversity in Bangladesh and Cambodia, while time
22 committed to domestic work is positively associated with diverse diets among women in
23 Cambodia, in Ghana (poorer households), and in Nepal (Komatsu et al., 2018). The authors
24 suggest that the positive association between domestic/care tasks and more diverse diets could
25 be a result of “staying close to the pot”. Studying the time allocation to leisure activities and

1 nutritional outcomes, Seymour et al., (2019) investigated the association of women's time
2 poverty and household nutrition in Bangladesh and found that women's time poverty (defined
3 as allocating less than 50 per cent of median time on leisure and self-care related activities) is
4 not significant in its association with household nutritional outcomes. Indeed, time-poor
5 women have relatively better nutritional outcomes.

6 The paradox seen in this strand of literature is that although female agricultural economic
7 time use suggests better nutritional outcomes through the increase in and control of incomes,
8 benefits can be outweighed by increasing time spent in strenuous physical activities leading to
9 greater energy expenditure (Nichols, 2016) and sociocultural norms entrenched in
10 intrahousehold negotiations can limit a woman's use of her monetary and time resources
11 (Agarwal, 1997; Bittman et al., 2003). However, Sangwan & Kumar (2021) and van den Bold
12 et al., (2021) finds no deleterious effects resulting from increasing agricultural time on
13 nutritional status. Their conclusions may be due to the small additional time spent in
14 agriculture following the interventions reported in their studies. Moreover, women may
15 regard improvements in household food security and income as beneficial even though such
16 involves trade-offs to their own well-being (Kabeer, 2001).

17 Reviews undertaken by FAO et al., (2020) and Johnston et al., (2018) show that the
18 information about men's time use is often less researched. Despite the significant focus on
19 women's time use and nutritional outcomes, women more than men continue to be
20 malnourished in rural areas of LMICs where most people depend on agriculture for their
21 livelihoods.

22 **2.2 Intrahousehold time allocation and nutritional externalities**

23 Intrahousehold externalities affect individual wellbeing (Basu, Narayan, & Ravallion, 2001)
24 but there are few pieces of empirical evidence assessing the relationship between men's time

1 allocation and women's nutritional outcomes and vice versa. Fleary & Joseph (2022) using
2 APIM to analyse data from the United States show interdependencies in health literacy, time
3 use and dietary behaviours between parents and adolescents. In the development literature,
4 intrahousehold externalities are largely streamlined to maternal time use and its consequences
5 on women's caring responsibilities for children's nutrition (Fadare et al., 2019; Ruel &
6 Alderman, 2013). Such focus on maternal time use and child nutrition is based on the
7 established linkages between the wellbeing of mother and child. However, in the face of
8 deprivation, gendered pay gaps and ownership of assets, the maintenance of adequate
9 nutrition among the poor and the very poor may lie in interdependencies between men and
10 women within the household (Rao et al., 2017). Such interdependent view has largely been
11 ignored, and women tend to have been targeted individually.

12 **3 Study area and data collection**

13 **3.1 Study area**

14 The secondary data used for analysis were collected in Jogulamba Gadwal District, south of
15 Telangana State in India. The district has 20 per cent scheduled castes, 1.5 per cent scheduled
16 tribes³ and more than three-quarters of its 609,990-population scattered in 255 rural villages.
17 About 60 per cent of its total land area is cultivated for food and cash crops, often on small
18 and marginal plots. Due to a substantial increase in the amount of monsoon rainfall and the
19 adoption of irrigation facilities in recent years, the semi-arid climate is increasingly turning
20 favourable to agricultural production (Government of Telangana, 2019).

21 **Figure 1 here**

³ "Schedule" refers to schedules in the Indian constitution identifying socially and economically deprived/marginalized caste groups and tribal (indigenous) groups as being entitled to affirmative actions in education, employment and development programs (Lelah Dushkin, 1967).

1 State government reports show a gradual decline in poverty in these areas; between 2014 and
2 2020, per capita annual income (adjusted for inflation) rose more than 10 per cent to 69,113
3 Rupees, equivalent to USD 1,100⁴. Rural income growth has been driven by agricultural
4 production expansion and participation in the Mahatma Gandhi National Rural Employment
5 Guarantee Act (MGNREGA) social welfare program. Data show that about 20 per cent of the
6 Jogulamba district population participates in MGNREGA and despite the mixed impacts⁵ of
7 the MGNREGA on agriculture in the area, agriculture and allied (crops, livestock, fisheries,
8 forestry) sector contribution to the overall product output rose to 21 per cent in 2021
9 (Government of Telangana, 2021). The growth can be attributed, in part, to other government
10 interventions in the form of inputs support, land redistribution, irrigation, and insurance
11 schemes.

12 Despite per-capita income increased in the study area over the last decade, malnutrition within
13 the population has remained high, especially among women. Figures from the Indian National
14 Family Health Survey show that 22 per cent of women and 17 per cent of men are
15 underweight (BMI < 18.5kg/m²) in rural Telangana in 2019-2020⁶; this is a decline from 29
16 per cent among women and 25 per cent among men in 2015-2016. The current prevalence of
17 anaemia among women is 58 per cent, up from 57 per cent in 2015-2016 (Christopher et al.,
18 2021; Ministry of Health and Family Welfare; Government of India, 2020). In comparison to
19 the other States in India, the high malnutrition rate is linked to the large number of scheduled
20 castes and scheduled tribes in Telangana.

⁴ 1 USD averaged 62.78 Indian Rupees in 2014 (Reserve Bank of India, 2020).

⁵ The MGNREGA has led to an increase in agricultural wages and a subsequent tightening of the agricultural labour market. In some instances, this agricultural labour shortage has been linked to shrinking farm plots in places where mechanization of farm work is elusive (Reddy et al., 2014).

⁶ Data was collected prior to the COVID-19 pandemic.

1 Further, the patterns of time use in this region show rural men and women commit over eight
2 hours to work-related activities daily (Government of India, 2020). There are however
3 substantial gender disparities albeit to a lesser degree compared to the rest of India: 55.7 per
4 cent of rural Telangana women participate in paid work, a figure three times the national
5 average (Government of India, 2020). Using data from the Time Use Survey-2019, Figure 2
6 shows the allocation of time among males and females living in rural areas of Telangana state.
7 Compared to men, women allocate on average 225 minutes more per day to care, domestic
8 and volunteer work, 158 minutes less to employment and production of goods for own use
9 and tend to spend on average 29 minutes more in work-related activities daily than men – the
10 time they seem to reallocate from socializing, self-care, and maintenance activities.

11 **Figure 2 here**

12 **3.2 Data collection**

13 *3.2.1 Survey*

14 The secondary dataset used in this paper is described in Zanello et al., (2020). Twenty
15 households were randomly selected after the households in the area had been stratified by
16 their ownership of irrigation infrastructure and the size of their landholdings⁷. In each
17 household, an economically active man and woman, aged between 16-64 years old took part
18 in the study. All households were employed primarily in crop production; eighteen
19 households cultivated their own land and two were sharecroppers. They cultivated
20 predominantly rice, cotton, yam, chillies, and groundnuts. Respondents were visited daily for
21 four non-consecutive weeks during June – November 2018, corresponding to each of the four

⁷ Data were collected in two communities – one using rainfed agriculture and one using irrigation infrastructure. The stratification of the sample households by landholding size control for differences in socio-economic characteristics across the two agricultural systems.

1 agricultural seasons of land preparation, sowing, land maintenance, and harvest when Kharif⁸
2 crops are cultivated.

3 At the beginning of the fieldwork, individuals self-reported information on their own health,
4 and anthropometric measurements of height and weight were taken. All the questionnaires
5 administered to respondents were translated into Telugu, the local language. The survey was
6 carried out by enumerators living in the same district and who spoke the local language.
7 Information on household characteristics was collected from the household head. In addition,
8 individual food intake data were collected daily based on a 24-hour recall throughout the four
9 weeks (Gibson & Ferguson, 2008). During the daily visits, enumerators also collected time
10 use information at one hour-intervals based on 24-hour recall.

11 3.2.2 *Accelerometers*

12 In addition to questionnaires administered daily, respondents were invited to wear an
13 accelerometer device throughout the length of the data collection. Accelerometers are portable
14 motion sensor devices used in the collection of objective physical activity data in free-living
15 populations (Troiano et al., 2014; Zanello et al., 2019). Raw 30Hz⁹ movement data were
16 collected using research-grade, tri-axial Actigraph GT3X+ accelerometers worn on the waist
17 by respondents during awake hours of 5 am – 11 pm. The movement data collected from
18 accelerometers were converted into energy expenditure (in kilocalories) using validated
19 algorithms (Freedson et al., 1998). Time use data collected using questionnaires were
20 matched with energy expenditure data derived from accelerometers to determine activity-
21 specific energy expenditure (Zanello et al., 2019).

⁸ In India, kharif crops are monsoon crops such as rice, maize, sugarcane, groundnut planted in July and harvested around October. Rabi crops are winter crops such as wheat, barley, carrot, chickpea planted in November and harvested around April and May.

⁹ 1Hz (Hertz) is one cycle per second.

1 While accelerometers provide an effective tool to capture energy expenditure in a free-living
2 population, they are not without limitations. Most importantly, as they capture movements,
3 they do not capture the additional effort involved in carrying weight or activities performed
4 while stationary (Lee & Shiroma, 2014). These limitations are particularly relevant for
5 activities typically categorized as domestic or caregiving, potentially leading to a greater
6 underestimation of energy expenditure in women compared to men (Shiroma et al., 2016). A
7 common challenge in physical activity research using accelerometers, especially in studies
8 among free-living populations, is participants not wearing the devices consistently (Troiano et
9 al., 2014). However, in our study, compliance with wearing the accelerometer was high. A
10 full day accelerometer wear rate was between 94-97% among the respondents. No sample
11 attrition was recorded during the four weeks of data collection.

12 The unique dataset used in this study therefore combines information on individual and
13 household sociodemographic characteristics, and individual data on food intake, time use, and
14 physical activity.

15 **4 Empirical Methods**

16 **4.1 Independent variables**

17 The main independent variables used in this study were time use variables measured as the
18 number of minutes allocated to each of economic, domestic and leisure activities (Moser,
19 1989). Every recorded activity in the hourly time use data was identified as either the primary
20 or secondary activity to ensure that typical secondary activities such as leisure and childcare
21 are also considered (Ironmonger, 2005). In cases where no secondary activities were recorded,
22 a weight of 1 was assigned to the hourly observation. A weight of 0.6 was assigned to primary
23 and weight of 0.4 was assigned to secondary activity, where respondents reported they carried

1 out simultaneous activities (Picchioni et al., 2020). We aggregate each of economic, domestic
2 and leisure time use data from hourly to day-level. Economic time use includes time spent in
3 agricultural activities such as crop and livestock production, forest produce collection and
4 related travel. Non-agricultural economic activities are salaried employment, non-farm wage
5 employment in construction and public work schemes, business, petty trading, and
6 professional development training. Domestic and care provision time use include household
7 maintenance and chores, food management, caring for children, elderly, sick and disabled.
8 Leisure time use includes time allocated to socializing and personal care.

9 **4.2 Dependent variables**

10 A set of three dependent variables are used in the analysis to capture the association between
11 time allocation and own and partner's calorie intake adequacy: Physical Activity Level
12 (PAL), Total Individual Energy Intake (EI), and Calorie Adequacy Ratio (CAR). We examine
13 the associations of time allocation with PAL, EI and CAR outcomes.

14 *4.2.1 Physical Activity Level (PAL)*

15 Physical Activity Level (PAL) is a measure of the intensity of physical activity over a day (or
16 other time period). To calculate individual PAL, raw 60-second epoch length physical activity
17 data collected from accelerometers were converted to Activity Energy Expenditure (AEE) in
18 kilocalories using a validated algorithm (Freedson et al., 1998). PAL is the ratio of Total
19 Energy Expenditure (TEE) to Basal Metabolic Rate (BMR), where TEE is the sum of BMR
20 (energy required to maintain vital physiological processes in the body) and AEE¹⁰. We
21 compute the BMR using the Harris-Benedict equation (Harris & Benedict, 1918). We use

¹⁰ TEE is the sum of BMR, AEE, and Thermal Effect of Feeding (TEF). TEF is energy required for metabolism, but TEF data is not available for this study. However, we assume the effect of this limitation to be minimal, since TEF accounts for only about 5-10 per cent of TEE (FAO, 2001).

1 PAL as the outcome variable instead of AEE because PAL controls for individual
2 anthropometric differences, allowing for comparisons across different age, gender, and BMI
3 groups. PAL values of 1.40-1.69 reflects sedentary or light activities, 1.70-1.99 moderate
4 activity and >2.00 indicates vigorous activity in free-living population. PAL has been used to
5 model energy expenditure among free-living populations (Friedman et al., 2021; Picchioni et
6 al., 2020; Srinivasan et al., 2020).

7 *4.2.2 Total Individual Energy Intake (EI)*

8 Total Individual Energy Intake (EI) is the total dietary energy reportedly consumed by
9 individual respondents in the last 24 hours. It captures the calorie (kcal) equivalent of food
10 and beverages per-adult day energy consumption (FAO, 2003). We use individual's food
11 intake data recorded through a 24-hour recall to compute the caloric values. The Indian Food
12 Composition tables were used to determine the calorie content of local recipes (Bowen et al.,
13 2011). The United States' National Nutrient Database for Standard Reference was used for
14 calorie conversion of ultra-processed foods (U.S. Department of Agriculture, 2019).
15 However, while EI captures caloric availability, the nutritional components of the food, and
16 the quality of diets cannot be ascertained.

17 *4.2.3 Calorie Adequacy Ratio (CAR)*

18 We use Calorie Adequacy Ratio (CAR) as a measure of nutritional outcomes. CAR is a metric
19 of energy balance which quantifies the overall dietary energy adequacy of an individual based
20 on the ratio of energy intake to energy expenditure (Randolph et al., 1991). We compute CAR
21 as the ratio of energy intake (EI) relative to total energy expenditure (TEE). An individual
22 whose CAR is equal to 1 is classified as energy balanced, a CAR below 1 is classified as
23 being energy deficient, and a CAR value above 1 indicates that the individual is in energy
24 surplus for a given day (FAO, 2001). The CAR as an indicator of nutritional outcomes allows

1 to measure individual energy intake adequacy. However, its focus on calories prevents
2 measuring the adequacy of the other nutrients necessary for a diverse diet. A person with a
3 CAR equal or above 1 may be deficient in essential nutrients. The description of all dependent
4 and independent variables used in the analysis (including intermediate variables) is presented
5 in Table 1.

6 **Table 1 here**

7 **4.3 Empirical strategy**

8 *4.3.1 The Actor-Partner Interdependence Model (APIM)*

9 The Actor-Partner Interdependence Model (APIM) explains dyadic relationships by
10 incorporating the concept of non-independence between two linked individuals with the
11 statistical methods to test such interdependence (Cook & Kenny, 2005). The APIM postulates
12 that own (actor), and partner's characteristics simultaneously influence the outcomes of both
13 individuals. This methodological approach assumes correlations in the characteristics and
14 outcomes of individuals within the same unit (for example, household). Conventional
15 statistical procedures assume independent observations but ignoring nonindependence of
16 observations between linked individuals will likely lead to biased statistical estimates (Cook
17 & Kenny, 2005). Non-independence in the observations of two linked persons may arise as a
18 result of *common fate*, *mutual influence* and *partner effects* (Kenny & Cook, 1999). APIM
19 focuses on modelling the interdependence between two individuals through partner effects.
20 Partner effects measure the bi-directional influence of one person on the other member of the
21 dyad. This contrasts intrahousehold behaviour theories that posits that individual outcomes
22 are determined either by individual preferences or by altruism (Fafchamps & Quisumbing,
23 2007). APIM approach has been used to study dyadic relationships, for example, in the

1 analysis of health behaviors in parent-adolescent dyads (Fleary & Joseph, 2022), work
2 division, communication, and couples' relationship satisfaction (Carlson et al., 2020). APIM
3 is used in this study to predict the influence that time allocation of spouses has on own and
4 partner's PAL, EI, and CAR outcomes. The household is treated as the unit of analysis.

5 **Figure 3 here**

6 Figure 3: Path depiction of the APIM model (Adapted from Kenny et al., 2006)

7 Notes: X_m = independent variable of the male, X_f = independent variable of the female, Y_m =
8 outcome variable of the male, Y_f = outcome variable of the female, β_m = male own (actor)
9 effects, ρ_m = male partner effects, β_f = female own (actor) effects, ρ_f = female partner effects,
10 E_1 and E_2 = error term.

11 We assess own (intrapersonal) and partner (interpersonal) effects of time use on dependent
12 variables of PAL, EI and CAR using the APIM for dyadic data depicted in Figure 4 (Cook &
13 Kenny, 2005). To treat individuals as nested within a dyad, we use the gender of each
14 respondent as the distinguishing variable within couples – and to capture role-specificity of
15 individuals. This differentiation allows for estimating the main components of the APIM: own
16 effects - β_m , β_f and partner effects - ρ_m , ρ_f , by using the main independent variables - X_m , X_f ;
17 and the dependent variables - Y_m and Y_f . Own effects (β_m , β_f) capture the association between
18 own independent variables and dependent variables (X_m ; and Y_m ; X_f and Y_f for male and
19 female respectively), while partner effects (ρ_m , ρ_f) capture the association between own
20 independent variables and partner's dependent variable (X_m and Y_f ; X_f and Y_m). E_1 and E_2
21 control for the correlation within couples. Interdependency between couple occurs when the
22 partner effects ρ_m , ρ_f are significantly associated with the dependent variables (Kenny et al.,
23 2006).

1 In the APIM model, to estimate own (β_m, β_f) and partner (ρ_m, ρ_f) effects of the time use
2 variables on the dependent variables of PAL, EI and CAR, we arrange the dyadic data in a
3 pairwise structure as shown in Table 2¹¹. Each row of data includes the household identifier,
4 the gender of the individual, the outcome variable for the individual and the characteristic(s)
5 of the individual. Additionally, the last two columns of Table-2 include the individual
6 characteristics and partner characteristics each multiplied by a dummy variable Z which is
7 equal to 0 for “own” and equal to 1 for “partner”. Arranging the data in this way yields
8 equations where the outcome for each individual is a function of the individual’s
9 characteristics and the partner’s characteristics.

10

11

Table 2 here

12 *4.3.2 Own and partner effects - couple composition, context, and the endogeneity of time use*
13 *variables.*

14 Ordinary least squares, structural equation modelling and multilevel modelling can be used in
15 the analysis of the APIM. We apply a multilevel model (MLM) to analyse the APIM
16 framework. This allows for the simultaneous estimation of hierarchies in the nested data- two
17 individuals (level-1) nested in a household (level-2) – whilst accounting for the inherent
18 nonindependence within each couple. To obtain the actor and partner effects by male and
19 female gender, the random two-intercepts model for MLM using the restricted maximum
20 likelihood method (Kenny et al., 2006; Rabe-Hesketh & Skrondal, 2012; Raudenbush et al.,
21 1995) estimates fifteen separate panel equations with separate observations for each day of the

¹¹ For simplicity, the illustration in Table-2 includes only one characteristic (explanatory variable) for each individual. The analysis can be extended to cases where there are several characteristics associated with each individual (e.g., time spent in different types of activities).

1 form $Y_{ijt} \in \{PAL_{ijt}, EI_{ijt}, CAR_{ijt}, CAR_{ijt} > 1, CAR_{ijt} < 1\}$ and $k = \{\text{Economic, Domestic and}$
 2 $\text{Leisure time uses}\}$:

$$3 \quad Y_{ijt} = \alpha_m^k \delta_i + \alpha_f^k (1 - \delta_i) + \beta_m^k X_{mjt}^k \delta_i + \beta_f^k X_{fjt}^k (1 - \delta_i) +$$

$$4 \quad \rho_m^k X_{mjt}^k (1 - \delta_i) + \rho_f^k X_{fjt}^k \delta_i + \theta_j^k \bar{X}_j^k + \omega^k \mathbf{I}_j + \sigma^k \mathbf{H}_j + \gamma^k \mathbf{C}_t + \tau^k \mathbf{Z}_s + \varepsilon_{ijt}^k \quad (1)$$

5 where i is the person (subscript $m = \text{male}, f = \text{female}$), j is household and t is day of the week;
 6 male α_m and female α_f intercepts; δ_i indicates that the person is male, female is $(1 - \delta_i)$; $X_{mjt}^k \delta_i$
 7 is the time spent in activities type k by the male in the j th household in t^{th} period (day);
 8 $X_{fjt}^k (1 - \delta_i)$ is the time spent in activities type k by the female in the j th household in t^{th} period
 9 (day); $X_{mjt}^k (1 - \delta_i)$ is the time spent in activities type k by the male partner in the j th household
 10 in t^{th} period (day); $X_{fjt}^k \delta_i$ is the time spent in activities type k by the female partner in the j th
 11 household in t^{th} period (day); \bar{X}_j^k is the mean of couple time use; \mathbf{I}_j is a vector of couple-mean
 12 centred variables of age and literacy¹², \mathbf{H}_j is vector of household socio-demographic
 13 characteristics such as irrigation system, size of cultivated land, household composition and
 14 assets index, and controls such as accelerometer wear, self-reported health, caste; \mathbf{C}_t is day
 15 dummies; \mathbf{Z} is seasonal (land preparation, sowing, land maintenance, and harvest) dummies;
 16 and the error term is $\varepsilon_{ijt} = \zeta_j + \mu_{ij}$ where ζ_j is household component, and individual-specific
 17 component μ_{ij}

18 The composition of groups, their contexts and the endogeneity of variables are likely sources
 19 of bias in multilevel analysis of APIM. For instance in our analysis, if higher couple literacy
 20 is associated with higher CAR for household j , comparing own and partner effects among
 21 couples is confounded by higher estimates among more literate couples (Bingenheimer &

¹² Couple mean centering of age and literacy was obtained by subtracting the household mean from individual observation.

1 Raudenbush, 2004; Rabe-Hesketh & Skrondal, 2012). This confounding by average
2 household level characteristics is referred to as compositional effects (Duncan et al., 1998).
3 We address compositional effects by including couple-mean centred variables of age and
4 literacy in equation 1 (Rabe-Hesketh & Skrondal, 2012).

5 In addition to bias that may be introduced by compositional effects, individual's patterns of
6 time use is known to correlate with unobserved household-level characteristics such as
7 sociocultural norms, resulting in level-2 endogeneity (Kevane & Wydick, 2001). We used the
8 Mundlak or "including-the-group-means approach" to address level-2 endogeneity of the time
9 use variables (Mundlak, 1978). This was done by including the means of couple time use
10 variables in equation 1. The Mundlak approach results in own and partner time use effects
11 that captures pure within-couple variation, which is unaffected by level-2 endogeneity.

12 Further, to ascertain the exogeneity of the within-couple time use estimates, we conduct post
13 regression tests of equal between and within time use effects (Rabe-Hesketh & Skrondal,
14 2012). Results show that the within-couple effects are uncorrelated with the between couple
15 time use effects. However, the Mundlak approach can produce biased estimates due to other
16 omitted variables, and the effects of time-invariant variables may not be consistent, as the
17 within and between effects are estimated separately in equation 1 (Hanchane & Mostafa,
18 2012). This limitation is addressed by the instrumental variable or Hausman-Taylor (HT)
19 approach (Hausman & Taylor, 1981). The HT approach can consistently estimate models with
20 endogenous time-invariant variables and time-variant variables, to produce estimates which
21 are uncorrelated with the residuals. In equation 2, household-level (level 2) factors not
22 captured in the model could have influenced differently, the patterns of time use of females
23 and males (level 1). The HT method first estimates individual-level effects of the time-
24 varying variables. This estimation produces residuals which are then regressed on time-

1 invariant variables. Regressing the residuals on the exogenous variables produces between-
 2 household effects, which are uncorrelated with the time-varying individual-level variables.
 3 The produced between-household effects act as instrumental variables (Rabe-Hesketh &
 4 Skrondal, 2012). As such, using the HT approach requires independent variables to be
 5 classified into four kinds as: exogenous time-varying variables, endogenous time-varying
 6 variables, exogenous time-constant variables, and endogenous time-constant variables. These
 7 are presented in Table 3.

8 **Table 3 here**

9 In addition to this criterion, the number of exogenous time-varying variables must be equal or
 10 higher than the number of the endogenous time-constant variables. Both conditions are
 11 satisfied in equation 2, where we estimated nine¹³ separate regression models where each
 12 outcome variable PAL, EI and CAR depend on each set of economic, domestic and leisure
 13 activities of the form $Y_{2ijt} \in \{PAL_{ijt}, EI_{ijt}, CAR_{ijt}\}$ and $k = \{\text{Economic, Domestic and}$
 14 $\text{Leisure}\}$:

$$15 \quad Y_{2ijt} = (\beta_{2i} + \zeta_j) + \beta_{2m}^k X_{mjt}^{k, \text{end}} \delta_i + \beta_{2f}^k X_{fjt}^{k, \text{end}} (1 - \delta_i) +$$

$$16 \quad \rho_{2m}^k X_{mjt}^{k, \text{end}} (1 - \delta_i) + \rho_{2f}^k X_{fjt}^{k, \text{end}} \delta_i + \pi_{ij}^k P_{ijt} + \omega_2^k I_j + \tau_2^k Z_s + \gamma_2^k C_t + \varepsilon_{ijt}^k \quad (2)$$

17 where i is the person (subscript $m = \text{male}$, $f = \text{female}$), j is household and t is on day t ;
 18 superscript *end* indicates endogenous variables; subscript 2 here distinguishes equation 1 from
 19 equation 2; $(\beta_{2i} + \zeta_j)$ is the intercept; δ_i indicates that the person is male = 1, female = $(1 - \delta_i)$;
 20 $X_{mjt}^{k, \text{end}}$ is the time spent in activities type k by the male in the j th household in t^{th} period (day);
 21 $X_{fjt}^{k, \text{end}}$ is the time spent in activities type k by the female in the j th household in t^{th} period

¹³ CAR>1 and CAR<1 was dropped as dependent variables in equation 2 because of the largely insignificant effects produced in equation 1.

1 (day); $X_{mjt}^{k, \text{end}}(1 - \delta_i)$ is the time spent in activities type k by the male partner in the jth
 2 household in t^{h} period (day); $X_{fjt}^{k, \text{end}} \delta_i$ is the time spent in activities type k by the female
 3 partner in the jth household in t^{h} period (day); P_{ijt}^k is a vector of gender and literacy; I_j is a
 4 vector of household socio-demographic characteristics such as irrigation system, size of
 5 cultivated land, vector of household composition and assets index, and controls such as
 6 accelerometer wear, self-reported health, caste; Z is seasonal (land preparation, sowing, land
 7 maintenance, and harvest) dummies; C_t is daily dummies; error term = ε_{ijt} . Own and partner
 8 time use variables were designated as related to components in the random intercept ($\beta_{2i} + \zeta_j$)
 9 in equation 2. The regression analysis was carried out using the “xthtaylor” command in Stata
 10 software (Castellano et al., 2014; Hausman & Taylor, 1981; StataCorp, 2013). The other form
 11 of endogeneity in MLM is the level-1 endogeneity of level-1 covariates. For instance,
 12 individual preference for certain activities may influence the amount of time spent on such
 13 activity. However, level-1 endogeneity in MLM is not directly testable (Rabe-Hesketh &
 14 Skrondal, 2012). Post-regression estimates of the own and partner effects of each time use
 15 category were computed as the percentage change in dependent variable divided by the
 16 percentage change in the independent variable for equations 1 and 2.

17 **5 Results**

18 **5.1 Descriptive statistics**

19 Table 4 presents descriptive statistics of household-level characteristics. On average,
 20 households in our sample cultivate around 10 acres of land, which is greater than the 3 acres
 21 district average (Government of Telangana, 2021). There is however variability in the sample
 22 with 35 per cent being smallholders, 35 per cent medium and 30 per cent large farmers based

1 on classification of landholding by the Indian Ministry of Agriculture and Farmers Welfare¹⁴.
2 The average household size of 4.3 is slightly below the Indian national average of 4.6 people
3 (UNDESA, 2019), with the number of males slightly higher than the number of females. The
4 respondents belonged to the backward caste, while one household identifies as belonging to
5 the scheduled caste.

6 **Table 4 here**

7 Descriptive statistics of individual-level characteristics are reported in Table 5. An average
8 PAL value of 1.55¹⁵ suggests that men and women spend a significant amount of time
9 engaged in light and moderate-intensity activities. There are indications of calorie deficits
10 among survey participants. Energy intakes for males and females are below the Indian
11 recommended daily dietary allowance (RDA) of 2,730 and 2,230 kcal for moderately active
12 people (National Institute of Nutrition, 2011). On average, men have a higher energy intake
13 than women (158 kcals/day more). However, relative to their energy expenditure needs, men
14 also have higher energy shortfalls compared to women. This translates to 978,82 kcals and
15 636,87 kcals calorie deficits for males and females respectively. Using CAR values, about 57
16 per cent of male respondents have an average daily CAR value below 1, while about 86 per
17 cent of female respondents have an average daily CAR value below 1. The CAR values
18 indicate more women than men are experiencing undernutrition.

19 **Table 5 here**

20 On average, males were older than females, with mean ages of around 40 years and a
21 significant difference of 5.5 years. In terms of literacy, defined as the ability to read and write,

¹⁴ Smallholders < 4.94 acres, medium 4.94-9.88 acres and large farmers >12.35 acres.

¹⁵ PAL values are classified as sedentary or light (1.40-1.69), active or moderately active (1.70-1.99), and vigorous (2.00-2.40) in free-living populations (FAO, 2001).

1 a substantial gap was observed: 30% of males were literate, compared to only 5% of females,
2 marking a significant difference of 25 percentage points. When it came to daily activities,
3 males spent significantly less time on domestic and care activities, averaging 27 minutes per
4 day, while females spent considerably more time, averaging 206 minutes per day. Conversely,
5 males engaged more in economic activities, averaging 516 minutes per day, compared to
6 females, who averaged 420 minutes per day. Men also spend on average 72 minutes more per
7 day than women in leisure activities. Similar unequal pattern of intrahousehold work division
8 have been reported in developed countries (Bittman et al., 2003).

9 **5.2 Own and partner effects**

10 As explained in Section 4.3.2, we run fifteen separate regressions such that own and partners'
11 time spent in economic, domestic and care and leisure activity were regressed on the outcome
12 variables of PAL, EI, CAR, CAR<1 and CAR>=1. Table 5 reports an overview of own and
13 partner effects elasticities computed post-MLM analysis of the Mundlak approach in Equation
14 1. The own and partner elasticities were computed as the percentage change in the dependent
15 variable relative to the percentage change in independent variable (that is, time use) in
16 minutes. The time use coefficients can be interpreted as the effect on the dependent variable,
17 of a one minute change in the time devoted to an activity category. Full regression tables are
18 reported in Appendix B. The effect sizes in Table 6 are expressed in percentages.

19 **Table 6 here**

20 *5.2.1 Physical Activity Level (PAL)*

21 The highest PAL effect is observed in the time allocated to economic activities, followed by
22 domestic and care activities. Conversely, the smallest PAL effect is noted in leisure activities,
23 for both females and males. A one percent increase in the time allocated to economic work

1 leads to a ten percent and eight percent increase in own PAL for males and females,
2 respectively. Male economic time use is associated with a two percent reduction in female
3 PAL. This suggests that male economic time use has a positive partner effect by reducing the
4 female PAL. Regarding domestic and care activity, a negative association is observed
5 between time spent on domestic and care work and PAL for both males and females.
6 Although very small, male domestic and care time use has a positive partner effect on female
7 PAL, while female domestic and care time use has no significant partner effect on male PAL.
8 Furthermore, given that men spend considerably less time on domestic and care activities, the
9 equal PAL effect size observed for domestic and care activity among males and females is
10 notable. This may be attributed to men engaging in short duration but more energy-intensive
11 activities, in contrast to women performing longer duration but less energy-intensive tasks.
12 Leisure time use is inversely related to PAL, with a more pronounced effect among women
13 than men. A one percent increase in the time allocated to leisure is associated with a five
14 percent reduction in PAL for females and a four percent reduction for males. Additionally, a
15 one percent increase in male leisure time use increases female PAL by one percent.

16 *5.2.2 Total Individual Energy Intake (EI)*

17 No significant effects on energy intake were observed for males across all three activity
18 categories. Female energy intake appears to increase by six percent with each one percent
19 increase in time spent on economic activities. These patterns of intrahousehold food
20 allocation indicate that time spent in economic activities is a significant, but not the only,
21 determinant of intrahousehold food distribution. The significant partner effects observed in
22 domestic and care time use confirm couple interdependence in this category and in terms of
23 energy intake. We also observe contrasting partner effects on energy intake; male time spent
24 in economic activity is correlated with a decrease in the quantity of food calories consumed

1 by his spouse, whereas female time spent in economic work is positively associated with the
2 energy intake of males. Given that the coefficient estimate for male partner economic time use
3 (-0.07) is lower than that for female partner economic time use (0.06), the effects of female
4 time spent in economic work on the couple's energy adequacy is large enough to offset the
5 reduced calorie intake due to male partner effects. In other words, females and their spouses
6 benefit more when the females spend time in economic work. This finding is consistent with
7 studies indicating that women's participation in economic work improves not only their
8 nutritional outcomes but also those of other household members (Ruel et al., 2018).

9 We also observe that female energy intake declines with increasing time allocated to domestic
10 and care activity, in contrast to the increasing effect observed with time spent in economic
11 work. This finding contradicts the positive nutritional outcomes associated with domestic and
12 care work reported in a previous multicountry study (Komatsu et al., 2018). The male partner
13 EI effects of domestic and care activity indicate an increase in female EI as men engage in
14 domestic and care activities. This pathway is further validated by the EI effects observed in
15 economic activities, where male participation in economic activities is seen to negatively
16 influence female EI. For both females and males, there is no statistically significant
17 relationship between EI and leisure time use.

18 *5.2.3 Calorie Adequacy Ratio (CAR)*

19 Results show that a one per cent increase in male economic time use leads to a nine per cent
20 decrease in CAR, with no corresponding significant relationship observed for females. This
21 contradicts the observations from Section 5.2.2, where female economic time use significantly
22 predicted EI but not CAR, in contrast to the findings for men. Such an outcome underscores
23 the importance of accounting for energy requirements in nutrition assessments. Regarding
24 partner effects, we note a decrease in female CAR with an increase in male economic time

1 use: a one per cent increase in male economic activity participation results in a five per cent
2 decrease in female CAR.

3 The effect of female domestic and care time use on female CAR mirrors that of male
4 economic activities on male CAR. Similar to the observations for EI, female domestic and
5 care time use reduces female CAR, while economic activities tend to reduce male CAR. In
6 addition to own effects, female domestic and care time use is also negatively associated with
7 the CAR of their spouses. Our results reveal no significant effects of male domestic and care
8 time use on CAR, neither in their own nor in partner effects.

9 The association between leisure time use and CAR shows that for every one per cent time
10 spent in leisure, female CAR increases by five per cent. Corresponding own male effects are
11 not significant, and there are no partner effects of leisure on CAR for both males and females.

12 We decompose CAR into energy sufficient ($CAR \geq 1$) and energy deficient ($CAR < 1$) groups to
13 provide additional insights on intra-couple time allocation by their energy adequacy status.

14 Calorie adequacy tends to decrease with increasing own domestic and care time use among
15 energy-deficient females. Among females with a calorie adequacy ratio greater than 1, CAR
16 appears to increase with male partner domestic and care time use. There are no significant
17 effects observed in the relationship between mean couple time use and CAR among the
18 calorie- deficient and sufficient groups.

19 **Table 7 here**

20 **5.3 Robustness checks**

21 To assess the robustness of regression results to the different estimation methods of the MLM,
22 we compare magnitudes and significance values between coefficient estimates of the
23 restricted maximum likelihood (REML) and full information maximum likelihood (FIML) for

1 all models. Estimates are similar and our conclusions hold for both REML and FIML
2 parameter estimation methods. However, our preferred approach is the REML methods, as it
3 is more suitable with estimations of small sample sizes compared to the FIML (Peugh, 2010).
4 The regression tables and post regression elasticities tables of the FIML are presented in
5 Appendices C1 – C4. Equation 2 regression results are presented in Table 7. Post-regression
6 elasticity estimates of the Hausman-Taylor estimator are quite like the Mundlak approach
7 already described in subsections 5.2.1, 5.2.2 and 5.2.3, except for the insignificant female
8 partner domestic and care time use effect on CAR in the Mundlak approach and the
9 insignificant male partner leisure time use on PAL in the Hausman-Taylor approach.

10 **6 Discussion and conclusion**

11 While time use patterns of women have been hypothesized to be responsible for the
12 persistence of malnutrition among women; previous empirical studies have mainly examined
13 the effects of women's time use allocations on children's nutrition. We contribute to the
14 literature on time allocation and nutritional externalities by looking at own and partner effects
15 of intra-couple time allocation on nutritional outcomes among households in Jogulamba
16 district in rural Telangana, India. Women tend to spend more time in domestic and care
17 activities in addition to economic activities, while men predominantly allocate more time to
18 economic activities.

19 The main finding of our analysis is that interdependencies between men and women in
20 households have an important influence on nutritional outcomes. Nutritional outcomes for
21 individuals are not determined by their own characteristics and endowments alone but also by
22 their spouses' characteristics and endowments. Specifically, increase in economic time
23 allocation by males, and an increase in domestic and care time allocation by females tend to

1 diminish both their own and their partners' food intake and caloric adequacy. Increasing time
2 spent in economic work is linked to improved nutrition for females, whereas for males,
3 nutritional improvements are associated with engaging in less physically intensive tasks, such
4 as domestic and care work. Partner effects reveal a negative link between female caloric
5 adequacy and increase in male economic time use, and a similar negative association exists
6 between male caloric adequacy and increase in female domestic and care time use. We see
7 that the greater participation of females in economic activities is rewarded with better
8 nutritional improvements relative to when males participate in economic activities, which
9 suggests that opportunities for women to participate in economic activities has the potential to
10 lead to improved nutritional outcomes in rural households.

11 Our analysis, incorporating both physical activity level and energy intake, highlights levels of
12 physical activity play an important role in caloric (in-) adequacy outcomes. This result
13 underscores the significance of considering physical activity information in individual
14 nutrition (energy requirement) assessments in rural areas of LMICs.

15 **6.1 Implications of study findings for development interventions**

16 The understanding that resources managed by women often lead to better household
17 nutritional outcomes than when those same resources are controlled by men has been a
18 foundational premise for many development and agricultural interventions. This approach is
19 predicated on the belief that women, when empowered with resources, are more likely to
20 allocate them in ways that enhance the nutritional well-being of their households. This
21 concept has been influential in shaping strategies that specifically target women with the aim
22 of achieving improved nutrition across households (Ruel et al., 2018). This paper underscores
23 the contribution of female economic work to securing own nutrition and that of other
24 members in line with existing literature.

1 Our findings reinforce the critical role that female economic participation plays in securing
2 not only their own nutritional well-being but also that of other household members. This
3 aligns with existing literature which finds positive impact of women's economic activity on
4 household nutrition (Kabeer, 2001; Quisumbing & Maluccio, 2003). By engaging in
5 economic work, women contribute significantly to household's resources, which in turn can
6 be leveraged to improve the nutritional status of the entire family. However, our findings
7 suggest that women's empowerment programmes focusing solely on increasing women's
8 productive assets may not guarantee improvements in nutritional or other outcomes, as these
9 outcomes also depend on personal and partner time allocations. To enhance nutritional
10 outcomes, women should no longer be regarded as the *sole proprietor* of household nutrition.
11 Development interventions should extend beyond improving autonomy for women, as current
12 evidence indicates that women in rural LMICs are already experiencing burdens in terms of
13 time and energy expenditure. Intrinsicly, the trade-offs to nutrition resulting from women's
14 empowerment or those occurring through the process of rural transformation can be
15 minimized by encouraging cooperation between spouses, especially regarding intra-household
16 sharing of domestic work. Indeed, changes in the norms surrounding intra-household work
17 division – supporting a gender-equal distribution of economic and domestic work – are
18 necessary to alleviate nutritional insecurity among both women and men. However, as spousal
19 cooperation tends to vary across socio-demographic contexts, it is important to understand
20 these contexts to tailor policy interventions aimed at advancing women's
21 empowerment (Kabeer, 2010; Lecoutere & Wuyts, 2020; Ragasa et al., 2019; Spark et al.,
22 2021).

23 Regarding whether increasing women's economic labour will not be detrimental to women's
24 health, growing evidence from the feminisation of agriculture literature reveals that increasing
25 female employment opportunities, especially in the agricultural sector, has not always led to

1 female empowerment. This is due to social norms and gender-intensified constraints such as
2 lack of productive assets, lower pay, and higher unpaid work burdens among women relative
3 to men (Asadullah & Kambhampati, 2021; Da Corta & Venkateshwarlu, 1999; Supriya
4 Garikipati, 2006). Policymakers concerned about female empowerment should address these
5 constraints and ensure that increases in female economic work are accompanied by a
6 reduction in their domestic and care work burdens. The way in which work and food are
7 shared between spouses will likely embody "unequal interdependence", where women bear
8 higher labour burdens relative to men (Kabeer, 2001). Yet, paid economic work constitutes
9 the beginning of "the breaking of traditional social norms" for some women, especially in
10 countries like India where female agricultural employment alone accounts for 58 per cent of
11 the 17 per cent total female labour force participation (Banerjee, 1997; ILO, 2022).

12 **6.2 Limitations and further research**

13 The innovative methodology used to collect and triangulate multiple data streams is not
14 without shortcomings. The sample size cannot be considered representative of the country
15 where the data was collected but rather an exemplary case study. Our empirical analysis is
16 supported by simulation studies that have proven that fixed-effects estimates (unlike variance
17 components) and standard errors of the multilevel analysis are not necessarily biased as a
18 result of sample size limitation (Bell et al., 2014; Huang, 2018; Peugh, 2010). Nevertheless,
19 weak significance values should be interpreted with caution. Also, due to statistical software
20 limitations, we have not examined heterogeneities across households through cross-level
21 effects of household characteristics or seasonality that may mediate the level of spousal
22 interdependency observed in this study. For instance, whether individual characteristics, type
23 of work, or income levels moderate own and partner effects.

1 Further, household composition has been shown in earlier literature to determine the division
2 of domestic work within couples with small children (Lundberg, 1988). Indeed, households in
3 our sample are composed of more than the two individuals that were sampled. Even if we had
4 no data for the other household members, we controlled in our analysis for the presence of
5 other members by including household size in the vector of household characteristics as well
6 as included a vector of seasonality to control for seasonal changes in time allocation.

7 Food intake data are known to be subjected to under-reporting due to social desirability and
8 recall bias, particularly in terms of food consumed outside the home. Under-reporting bias in
9 our study may be larger for men than women concerning calories derived from alcohol
10 consumption and food consumed outside the home. Also, cultural aspects of intrahousehold
11 food sharing such as the order of food servings and the tendency to allocate more nutritious
12 meals to males – are not explicitly considered in this study due to data limitations. Given the
13 focus of this paper on calories, the indicators of nutritional outcomes are not comprehensive
14 to understand the nutrient adequacy and healthiness of diets. These are aspects that future
15 work may seek to improve upon.

16

1 **Information from human participants**

2 All personal information that would allow the identification of any person or person(s)

3 described in the article has been removed. The data used is contained in Zanello et al., (2020).

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